## **Balanced Allocations: A Refined Drift Theorem**

#### <u>Dimitrios Los</u><sup>1</sup>, Thomas Sauerwald<sup>1</sup>, John Sylvester<sup>2</sup>

<sup>1</sup>University of Cambridge, UK, <sup>2</sup>University of Liverpool, UK



Based on: "An Improved Drift Theorem for Balanced Allocations" (arXiv) & "Balanced Allocations with Heterogeneous Bins: The Power of Memory" (arXiv)

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## Outline

Balanced allocations (background and some highlights)

The exponential and hyperbolic cosine potential functions

The proof of the drift theorem

The refinement and its applications

Open problems

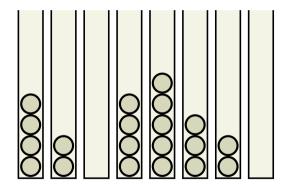
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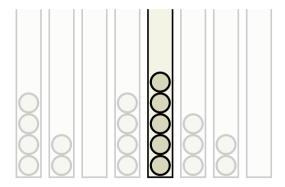
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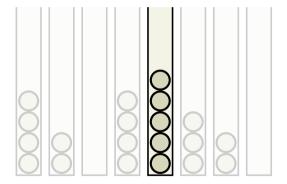
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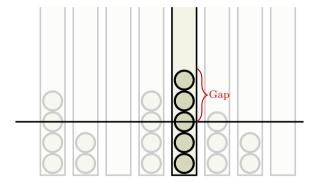
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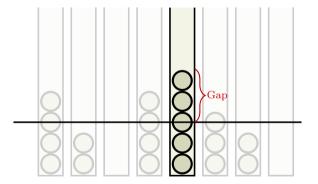
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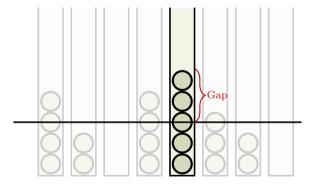
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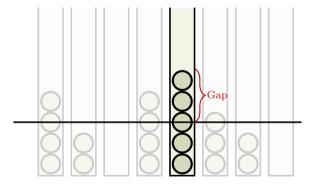
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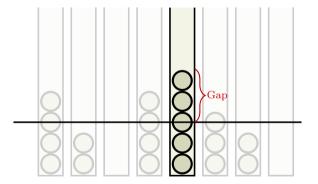
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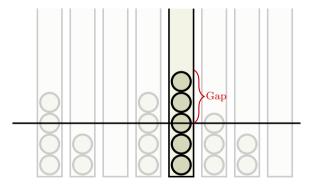
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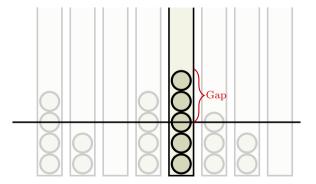
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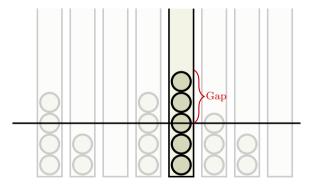
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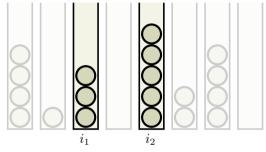
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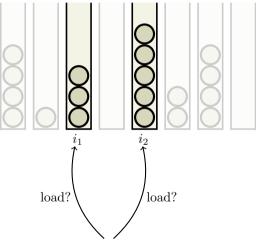
Question: Why variants and not vanilla TWO-CHOICE?

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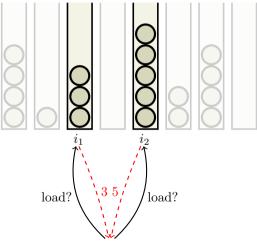
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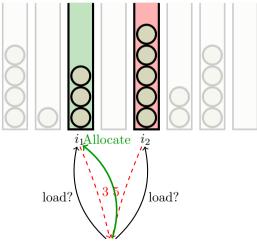
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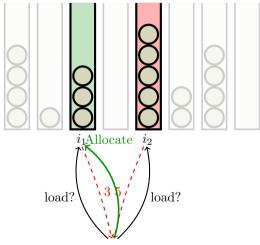
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We have shown that several natural selection rules are not optimal in various situations, but we have not identified any optimal rules. Identifying optimal rules in these situations would obviously be interesting, but appears to be difficult.

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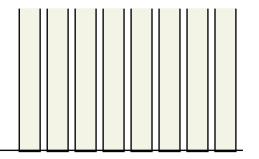
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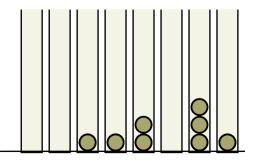
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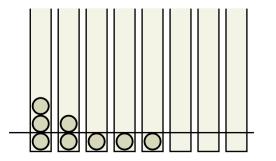
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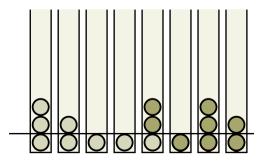
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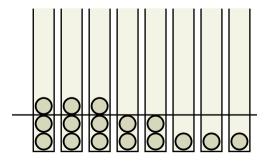
**Question:** Why choose a  $\beta < 1$ ?

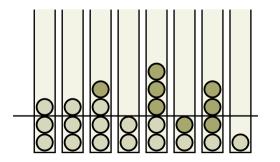


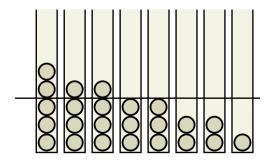


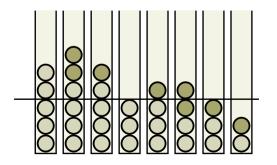


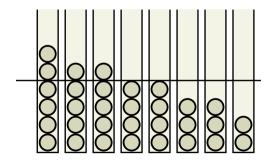


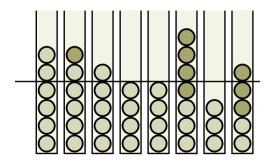


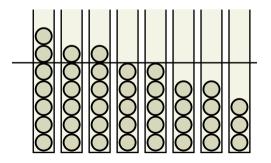






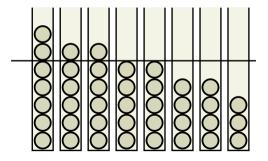




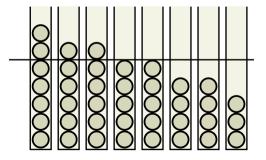


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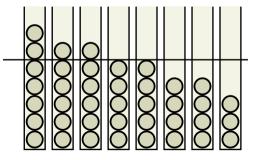


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#### Probability allocation vectors

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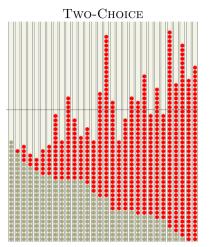
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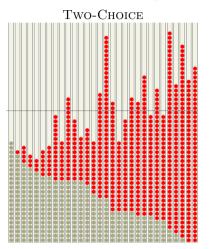
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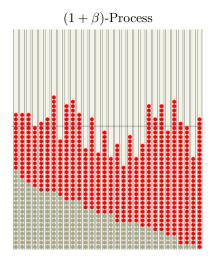
For  $(1 + \beta)$ -process,  $p_{(1+\beta)} = \left(\dots, \beta \cdot \frac{2i-1}{n^2} + (1-\beta) \cdot \frac{1}{n}, \dots\right).$ 



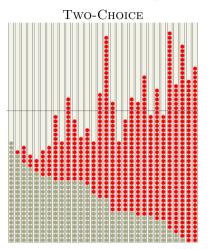
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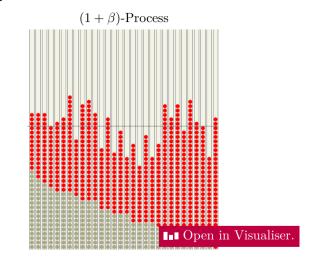


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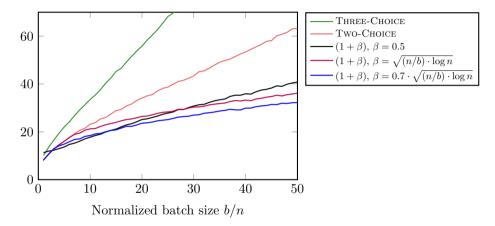
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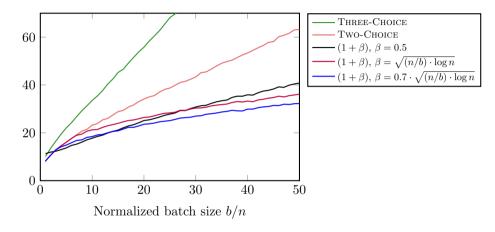
An example of a variant of Two-Choice



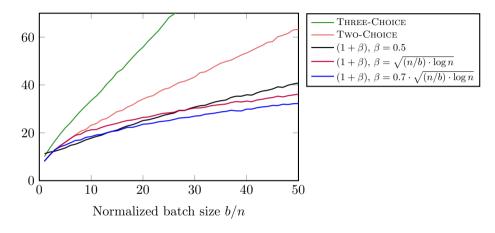
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 $^{12}$ 

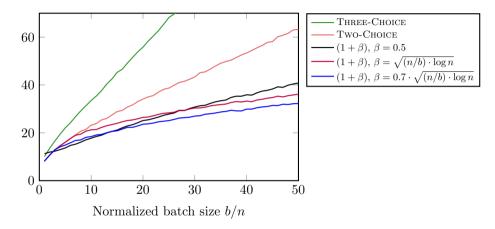




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# **Potential functions**

# Techniques for analyzing balanced allocations

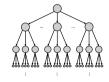
#### Layered induction

Witness trees



TWO-CHOICE, MEMORY

Poissonisation



TWO-CHOICE, parallel allocations

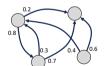
Markov chains

Graphical processes



TWO-CHOICE

Potential functions



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- 2-2				 	 
	-	-	-	 	 
				_	_
$\delta_1 \delta_1$	2				-

$$X_i \sim \mathsf{Poi}(\frac{m}{n})$$

 $Unweighted,\ time-independent$ 

Some weights, *b*-BATCHED, heterogeneous sampling

weights, b-BATCHED, outdated info, noise graphical, heterogeneous sampling

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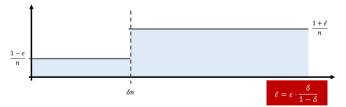
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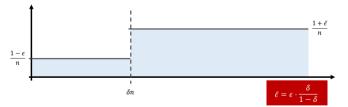


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Our main aim will be to derive the w.h.p.  $\mathcal{O}((\log n)/\epsilon)$  gap, for any  $\epsilon \in (0, 1)$ .

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Then, applying *Markov's inequality* we get that w.h.p.  $\mathbf{E}[\Phi^t] = \text{poly}(n)$ .

Let us fix a bin  $i \in [n]$ .

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Applies also to weights  $\mathcal{W}$  with unit expectation and finite MGF, i.e.,  $e^{\alpha \mathcal{W}} \leq 1 + \alpha + \alpha^2 \cdot S.$ 

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using the Taylor estimate  $e^z \le 1 + z + z^2$  for sufficiently small For bins with  $p_i = \frac{1-\epsilon}{n}$ ,

$$\begin{aligned} \mathbf{E} \text{ Let us fix a bin } i \in [n]. \text{ Then,} \\ \mathbf{E} \left[ \Phi_i^{t+1} \mid \Phi_i^t \right] &= p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)} \\ &\leq \Phi_i^t \cdot \left( p_i \cdot \left( 1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left( 1 - \alpha/n + \alpha^2/n^2 \right) \right) \\ &= \Phi_i^t \cdot \left( 1 + \alpha \cdot \left( p_i - \frac{1}{n} \right) + \alpha^2 \cdot \left( p_i + \frac{1 - p_i}{n} \right) \right), \\ \text{using the Taylor estimate } e^z \leq 1 + z + z^2 \text{ for sufficiently small } z. \end{aligned}$$

$$\begin{aligned} &= \text{ For bins with } p_i = \frac{1 - \epsilon}{n}, \text{ we have that } \left( \Delta \Phi_i^{t+1} \coloneqq \Phi_i^{t+1} - \Phi_i^t \right) \\ &= \mathbf{E} \left[ \Delta \Phi_i^{t+1} \mid \mathfrak{F} \right] \leq \Phi_i^t \cdot \left( -\frac{\alpha \epsilon}{n} + \mathcal{O} \left( \frac{\alpha^2}{n} \right) \right) \end{aligned}$$

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$$\textbf{For bins with } p_i = \frac{1 - \epsilon}{n}, \text{ we have that } (\Delta \Phi_i^{t+1} := \Phi_i^{t+1} - \Phi_i^t) \\ \textbf{E} \left[ \Delta \Phi_i^{t+1} \mid \mathfrak{F}^t \right] \leq \Phi_i^t \cdot \left( -\frac{\alpha \epsilon}{n} + \mathcal{O} \left( \frac{\alpha^2}{n} \right) \right) \rightsquigarrow \text{Good bin.} \end{split}$$

Let us fix a bin 
$$i \in [n]$$
. Then,
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$$\leq \Phi_{i}^{t} \cdot \left(p_{i} \cdot \left(1+\alpha \cdot (1-1/n)+\alpha^{2}\right) + (1-p_{i}) \cdot \left(1-\alpha/n+\alpha^{2}/n^{2}\right)\right)$$

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• Otherwise, for bins with  $p_i = \frac{1+\epsilon}{n}$ ,  $\mathbf{E}\left[\Delta \Phi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi_i^t \cdot \left(+\frac{\alpha \widetilde{\epsilon}}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right)$ 

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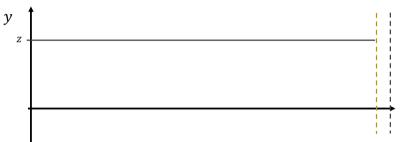
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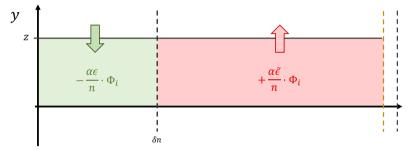
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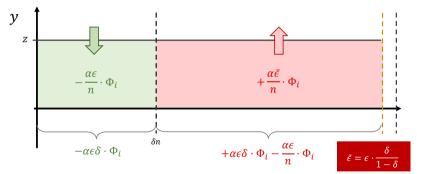
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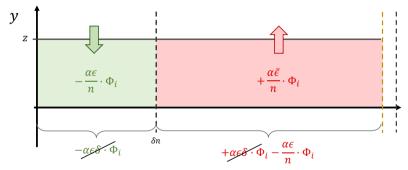
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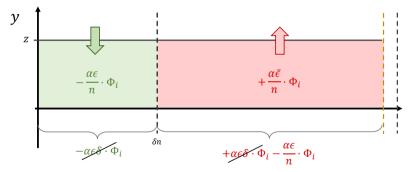


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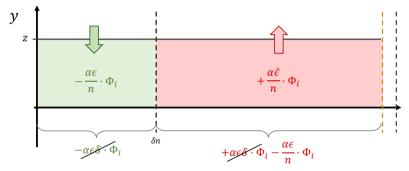
Consider a step  $t \ge 0$ , where all but one bins have the same load:



We could get only a *very small* decrease.

There could be too many overloaded bins.

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We could get only a very small decrease.  $\rightsquigarrow$  Gives  $\mathcal{O}(n \log n/\epsilon)$  bound on the gap.

Potential functions

The *hyperbolic cosine potential* [PTW15, Spe77] is defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} + e^{-\alpha \cdot (x_i^t - t/n)}$$

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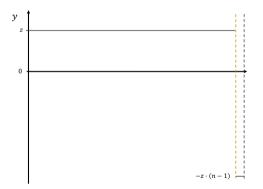
$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} + e^{-\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t} + e^{-\alpha \cdot y_i^t} + e$$

**Question:** Why does the *second term* help?

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Potential functions

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So, the dominant term is the *negative* of that in Φ.
 More specifically, if p<sub>i</sub> = 1+ϵ̃/n, then
 E [ΔΨ<sub>i</sub><sup>t+1</sup> | 𝔅<sup>t</sup>] ≤ Ψ<sub>i</sub><sup>t</sup> ⋅ (-αϵ̃/n + O(α<sup>2</sup>/n)) → Good bin.

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So, the dominant term is the *negative* of that in  $\Phi$ . More specifically, if  $p_i = \frac{1+\widetilde{\epsilon}}{n}$ , then  $\mathbf{E}\left[\Delta \Psi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Psi_i^t \cdot \left(-\frac{\alpha \widetilde{\epsilon}}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right) \rightsquigarrow \text{Good bin.}$ 

Otherwise, if  $p_i = \frac{1-\epsilon}{n}$   $\mathbf{E}\left[\Delta \Psi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Psi_i^t \cdot \left(+\frac{\alpha \epsilon}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right) \rightsquigarrow \text{Bad bin.}$ 

# Drift Theorem

#### Theorem ([PTW15, Section 2])

Consider any process with *non-decreasing* allocation vector p which is  $\epsilon$ -biased for some  $\epsilon \in (0, 1)$  and some constant  $\delta$ , in the setting with weights sampled from a distribution with finite MGF. Then, for  $\Gamma := \Gamma(\alpha)$  with  $\alpha := \Theta(\epsilon)$ , for any step  $t \ge 0$ ,

$$\mathbf{E}\left[\left|\Delta\Gamma^{t+1}\right| \,\mathfrak{F}^{t}\right] \leq -\Gamma^{t} \cdot \frac{\alpha\epsilon}{4n} + \operatorname{poly}(1/\epsilon),$$

and

 $\mathbf{E}\left[\,\Gamma^t\,\right] \le n \cdot \operatorname{poly}(1/\epsilon).$ 

## **Refined Drift Theorem**

#### Theorem ([LS22, Corollary 3.2])

Consider any process and a *probability vector* p being  $\epsilon$ -biased for some  $\epsilon \in (0, 1)$  and some constant  $\delta$ . Further assume that it satisfies for some K > 0 and for any  $t \ge 0$ ,

$$\mathbf{E}\left[\left|\Phi^{t+1}\right| \mathfrak{F}^{t}\right] \leq \sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left(1 + \left(p_{i} - \frac{1}{n}\right) \cdot \alpha + K \cdot \frac{\alpha^{2}}{n}\right),$$

and

$$\mathbf{E}\left[\left|\Psi^{t+1}\right| \mathfrak{F}^{t}\right] \leq \sum_{i=1}^{n} \Psi_{i}^{t} \cdot \left(1 + \left(\frac{1}{n} - p_{i}\right) \cdot \alpha + K \cdot \frac{\alpha^{2}}{n}\right).$$

Then, there exists a constant  $c := c(\delta) > 0$ , such that for  $\alpha \in \left(0, \min\left\{1, \frac{\epsilon\delta}{8K}\right\}\right)$ 

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot \left(1 - \frac{\alpha \epsilon \delta}{8n}\right) + c\alpha \epsilon,$$

and

$$\mathbf{E}\left[\,\Gamma^t\,\right] \le \frac{8c}{\delta} \cdot n.$$

## **Refined Drift Theorem**

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Consider any process and a probability vector p being  $\epsilon$ -biased for some  $\epsilon \in (0, 1)$  and some constant  $\delta$ . Further assume that it satisfies for some K > 0 and some R > 0, for any  $t \ge 0$ ,

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and

$$\mathbf{E}\left[\left|\Psi^{t+1}\right| \mathfrak{F}^{t}\right] \leq \sum_{i=1}^{n} \Psi_{i}^{t} \cdot \left(1 + \left(\frac{1}{n} - p_{i}\right) \cdot \frac{\mathbf{R}}{\mathbf{R}} \cdot \alpha + K \cdot \frac{\mathbf{R}}{n} \cdot \frac{\alpha^{2}}{n}\right)$$

Then, there exists a constant  $c := c(\delta) > 0$ , such that for  $\alpha \in \left(0, \min\left\{1, \frac{\epsilon\delta}{8K}\right\}\right)$ 

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot \mathbf{R} \cdot \left(1 - \frac{\alpha \epsilon \delta}{8n}\right) + \mathbf{R} \cdot c\alpha \epsilon,$$

and

$$\mathbf{E}\left[\,\Gamma^t\,\right] \le \frac{8c}{\delta} \cdot n.$$

Our goal is to show that:

$$\sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left( \alpha \cdot \left( p_{i} - \frac{1}{n} \right) + K \cdot \frac{\alpha^{2}}{n} \right) + \Psi_{i}^{t} \cdot \left( \alpha \cdot \left( \frac{1}{n} - p_{i} \right) + K \cdot \frac{\alpha^{2}}{n} \right)$$

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$$\leq -\frac{\alpha \epsilon}{4n} \cdot \Gamma^{t} + K \cdot \frac{\alpha^{2}}{n} \cdot \Gamma^{t} + \mathcal{O}(\alpha \epsilon)$$

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 for any  $\alpha \leq \frac{\epsilon}{8K}$ . (Key Observation 2)

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We have the following types of bins

$\operatorname{Set}$	Load	Index	$r_i$	Dominant Contribution
Good overloaded $\mathcal{G}_+$	$y_i \ge 0$	$i \leq \delta n$	$\frac{1-\epsilon}{n_{\sim}}$	$-\Phi_i \cdot \frac{\alpha \epsilon}{n \epsilon} + \Psi_i \cdot \frac{\alpha \epsilon}{n \epsilon}$
Bad overloaded $\mathcal{B}_+$	$y_i \ge 0$	$i > \delta n$	$\frac{1+\epsilon}{n_{\sim}}$	$+\Phi_i \cdot \frac{\alpha\epsilon}{n_{\sim}} - \Psi_i \cdot \frac{\alpha\epsilon}{n_{\sim}}$
Good underloaded $\mathcal{G}_{-}$	$y_i < 0$	$i > \delta n$	$\frac{1+\epsilon}{n}$	$+\Phi_i \cdot \frac{\alpha\epsilon}{n} - \Psi_i \cdot \frac{\alpha\epsilon}{n}$
Bad overloaded $\mathcal{B}_{-}$	$y_i < 0$	$i \leq \delta n$	$\frac{1-\epsilon}{n}$	$-\Phi_i \cdot \frac{lpha\epsilon}{n} + \Psi_i \cdot \frac{lpha\epsilon}{n}$

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$$\begin{split} \sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left( \alpha \cdot \left( p_{i} - \frac{1}{n} \right) + K \cdot \frac{\alpha^{2}}{n} \right) + \Psi_{i}^{t} \cdot \left( \alpha \cdot \left( \frac{1}{n} - p_{i} \right) + K \cdot \frac{\alpha^{2}}{n} \right) \\ & \leq -\frac{\alpha \epsilon}{4n} \cdot \Gamma^{t} + K \cdot \frac{\alpha^{2}}{n} \cdot \Gamma^{t} + \mathcal{O}(\alpha \epsilon) \\ & \leq -\frac{\alpha \epsilon}{8n} \cdot \Gamma^{t} + \mathcal{O}(\alpha \epsilon), \end{split}$$

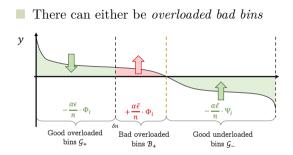
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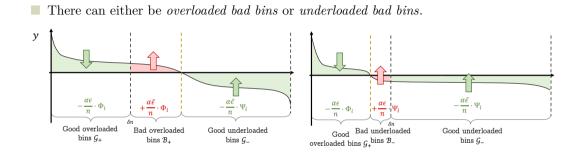
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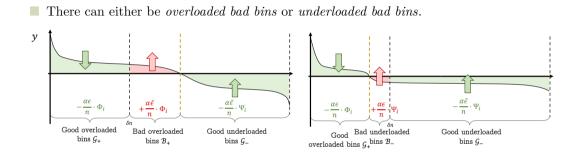
**Key Observation 3:** For overloaded bins  $\Psi_i^t \leq 1$  and for underloaded bins  $\Phi_i^t \leq 1$ ,  $\rightsquigarrow$  their contribution is  $\mathcal{O}(\alpha \epsilon)$ .

Potential functions

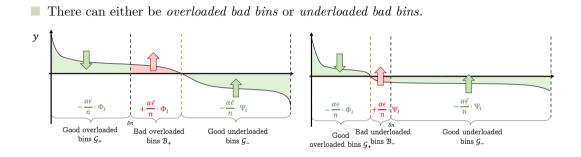


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**Key Observation 4:** The second case is symmetric to the first:  $\delta' = 1 - \delta$ ,  $\Phi' = \Psi$ ,  $\Psi' = \Phi$  and y' = -y.

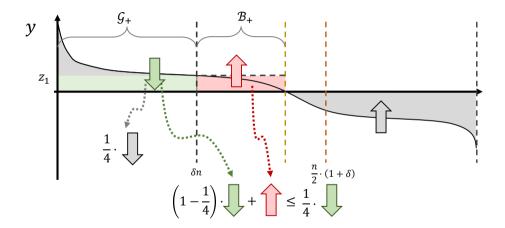


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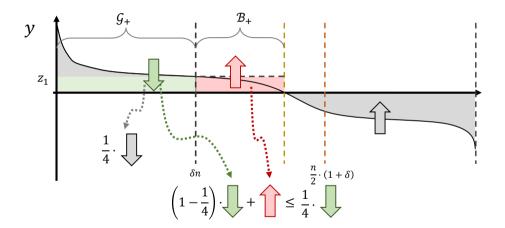
So we only consider Case A.

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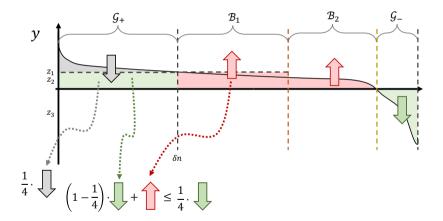


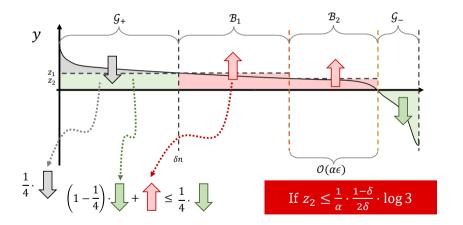
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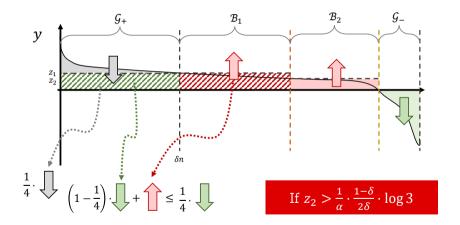


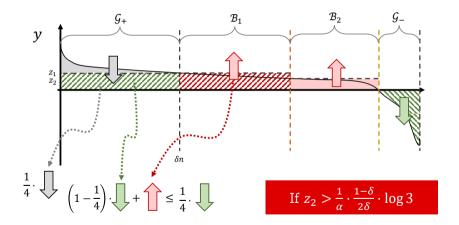
As with the exponential potential, we counteract the bad bins with a fraction of the decrease of the overloaded good bins. *All* underloaded bins are good.

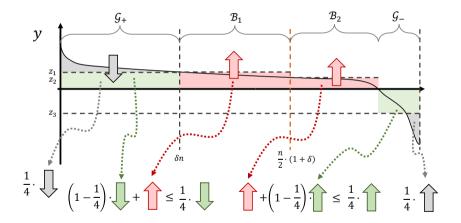
Potential functions











We used the following techniques:

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- 5. Use the decrease of the underload potential to counteract the increase of bad bins.

## The drift theorem

#### Theorem ([LS22, Corollary 3.2])

Consider any allocation process and a probability vector p being  $\epsilon$ -biased for some  $\epsilon \in (0, 1)$  and some constant  $\delta$ . Further assume that it satisfies for some K > 0 and some R > 0, for any  $t \ge 0$ ,

$$\mathbf{E}\left[\left|\Phi^{t+1}\right|\left|\mathfrak{F}^{t}\right|\right] \leq \sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left(1 + \left(p_{i} - \frac{1}{n}\right) \cdot R \cdot \alpha + K \cdot R \cdot \frac{\alpha^{2}}{n}\right)$$

and

$$\mathbb{E}\left[\left|\Psi^{t+1}\right|\left|\mathfrak{F}^{t}\right|\right] \leq \sum_{i=1}^{n} \Psi_{i}^{t} \cdot \left(1 + \left(\frac{1}{n} - p_{i}\right) \cdot R \cdot \alpha + K \cdot R \cdot \frac{\alpha^{2}}{n}\right)$$

Then, there exists a constant  $c := c(\delta) > 0$ , such that for  $\alpha \in \left(0, \min\left\{1, \frac{\epsilon \delta}{8K}\right\}\right)$ 

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot R \cdot \left(1 - \frac{\alpha \epsilon \delta}{8n}\right) + R \cdot c \alpha \epsilon,$$

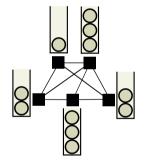
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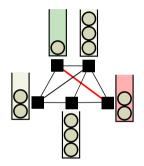
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# Applications

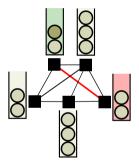
Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:



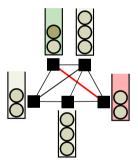
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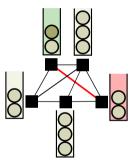


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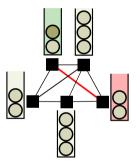
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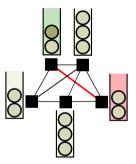
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## Example 1: The Graphical Setting

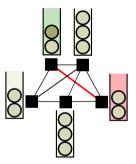
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For any *d*-regular graph with conductance ε, p<sup>t</sup> is majorized by an ε-biased probability vector. → gap is w.h.p. O(log n)/ε [PTW15].
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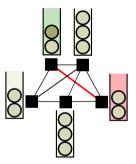


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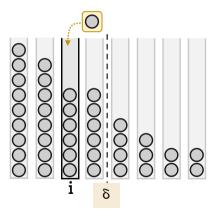
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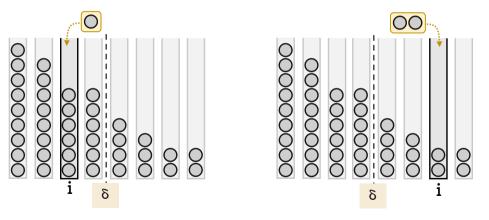
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For processes allocating *more than one* balls we have that:

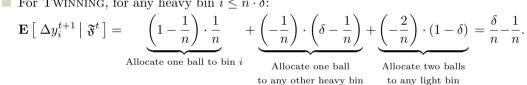
$$\begin{split} & \mathbf{E}\left[\left.\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right]\leq\Phi_{i}^{t}\cdot\left(\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right),\\ & \mathbf{E}\left[\left.\Delta\Psi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Psi_{i}^{t}\cdot\left(-\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right). \end{split}$$

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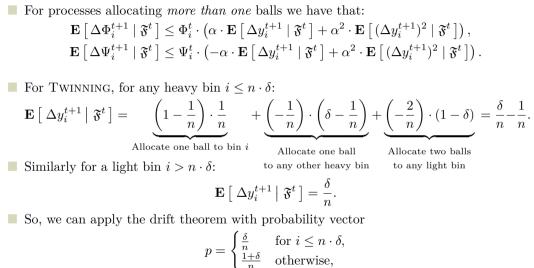
For TWINNING, for any heavy bin 
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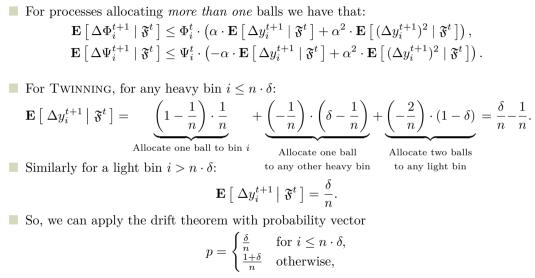
$$\mathbf{E} \left[ \Delta y_i^{t+1} \mid \mathfrak{F}^t \right] = \underbrace{\left( 1 - \frac{1}{n} \right) \cdot \frac{1}{n}}_{\text{Allocate one ball to bin } i} + \underbrace{\left( -\frac{1}{n} \right) \cdot \left( \delta - \frac{1}{n} \right)}_{\text{Allocate one ball to any other heavy bin}} + \underbrace{\left( -\frac{2}{n} \right) \cdot (1 - \delta)}_{\text{Allocate two balls to any light bin}}$$

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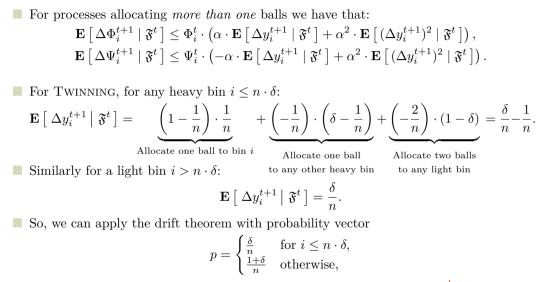
For processes allocating *more than one* balls we have that:  $\mathbf{E}\left[\Delta\Phi_{i}^{t+1} \mid \mathfrak{F}^{t}\right] < \Phi_{i}^{t} \cdot \left(\alpha \cdot \mathbf{E}\left[\Delta y_{i}^{t+1} \mid \mathfrak{F}^{t}\right] + \alpha^{2} \cdot \mathbf{E}\left[\left(\Delta y_{i}^{t+1}\right)^{2} \mid \mathfrak{F}^{t}\right]\right),$  $\mathbf{E}\left[\Delta\Psi_{i}^{t+1} \mid \mathfrak{F}^{t}\right] \leq \Psi_{i}^{t} \cdot \left(-\alpha \cdot \mathbf{E}\left[\Delta y_{i}^{t+1} \mid \mathfrak{F}^{t}\right] + \alpha^{2} \cdot \mathbf{E}\left[(\Delta y_{i}^{t+1})^{2} \mid \mathfrak{F}^{t}\right]\right).$ For TWINNING, for any heavy bin  $i \leq n \cdot \delta$ :  $\mathbf{E}\left[\left.\Delta y_i^{t+1} \right| \,\mathfrak{F}^t \right] = - \left(1 - \frac{1}{n}\right) \cdot \frac{1}{n} - + \left(-\frac{1}{n}\right) \cdot \left(\delta - \frac{1}{n}\right) + \left(-\frac{2}{n}\right) \cdot (1 - \delta) = \frac{\delta}{n} - \frac{1}{n}.$ Allocate one ball to bin iAllocate one ball Allocate two balls to any other heavy bin to any light bin Similarly for a light bin  $i > n \cdot \delta$ :  $\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\right|\left.\mathfrak{F}^{t}\right.\right]=\frac{\delta}{\pi}.$ 





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Applications



which is  $\epsilon$ -biased with  $\epsilon = 1 - \delta$ . So, by the *drift theorem*, we get an  $\mathcal{O}(\frac{\log n}{1-\delta})$  gap.

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- Again, applying the drift theorem gives w.h.p. an  $\mathcal{O}(\log n)$  upper bound on the gap.

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# Conclusions

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- It is *agnostic* of the balanced allocations setting.

Still many open problems, including:

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- Can we use a *single potential* to prove sublogarithmic bounds (e.g., the  $\log_2 \log n + \Theta(1)$  bound for TWO-CHOICE)?

#### Questions?

 $More \ visualisations: \ \tt dimitrioslos.com/wand-disc23$ 

## Bibliography I

- ▶ D. Alistarh, J. Aspnes, and R. Gelashvili, *Space-optimal majority in population protocols*, 29th Annual ACM-SIAM Symposium on Discrete Algorithms (SODA'18), SIAM, 2018, pp. 2221–2239.
- D. Alistarh, t. Brown, J. Kopinsky, J. Z. Li, and G. Nadiradze, *Distributionally linearizable data structures*, 30th Annual ACM Symposium on Parallel Algorithms and Architectures (SPAA'18), ACM, 2018, pp. 133–142.
- Y. Azar, A. Z. Broder, A. R. Karlin, M. Mitzenmacher, and E. Upfal, The ACM Paris Kanellakis Theory and Practice Award, 2020, https://www.acm.org/media-center/2021/may/technical-awards-2020.
- Y. Azar, A. Z. Broder, A. R. Karlin, and E. Upfal, *Balanced allocations*, SIAM J. Comput. **29** (1999), no. 1, 180–200. MR 1710347
- ▶ D. Alistarh, R. Gelashvili, and J. Rybicki, *Fast graphical population protocols*, 25th International Conference on Principles of Distributed Systems (OPODIS'21), vol. 217, Schloss Dagstuhl Leibniz-Zentrum für Informatik, 2021, pp. 14:1–14:18.

## **Bibliography II**

- D. Alistarh, J. Kopinsky, J. Li, and g. Nadiradze, *The power of choice in priority scheduling*, 36th Annual ACM-SIGOPT Principles of Distributed Computing (PODC'17), ACM, 2017, pp. 283–292.
- ▶ P. Berenbrink, A. Czumaj, M. Englert, T. Friedetzky, and L. Nagel, *Multiple-choice* balanced allocation in (almost) parallel, 16th International Workshop on Randomization and Computation (RANDOM'12), Springer-Verlag, 2012, pp. 411–422.
- P. Berenbrink, A. Czumaj, A. Steger, and B. Vöcking, Balanced allocations: the heavily loaded case, SIAM J. Comput. 35 (2006), no. 6, 1350–1385. MR 2217150
- M. Dahlin, Interpreting stale load information, IEEE Trans. Parallel Distributed Syst.11 (2000), no. 10, 1033–1047.
- P. Delgado, D. Didona, F. Dinu, and W. Zwaenepoel, Job-aware scheduling in eagle: Divide and stick to your probes, 7th ACM Symposium on Cloud Computing (SoCC'16), ACM, 2016, pp. 497–509.

## Bibliography III

- P. Delgado, F. Dinu, A. M. Kermarrec, and W. Zwaenepoel, *Hawk: Hybrid datacenter scheduling*, 2015 USENIX Annual Technical Conference (USENIX'15), USENIX, 2015, pp. 499–510.
- A. Gupta, R. Krishnaswamy, A. Kumar, and S. Singla, Online carpooling using expander decompositions, 40th IARCS Annual Conference on Foundations of Software Technology and Theoretical Computer Science (FSTTCS'20), vol. 182, Schloss Dagstuhl -Leibniz-Zentrum für Informatik, 2020, pp. 23:1–23:14.
- ▶ G. H. Gonnet, Expected length of the longest probe sequence in hash code searching, J. Assoc. Comput. Mach. 28 (1981), no. 2, 289–304. MR 612082
- M. Khelghatdoust and V. Gramoli, *Peacock: Probe-based scheduling of jobs by rotating between elastic queues*, 24th International Conference on Parallel and Distributed Computing (Euro-Par'18), vol. 11014, Springer, 2018, pp. 178–191.
- R. M. Karp, M. Luby, and F. Meyer auf der Heide, Efficient PRAM simulation on a distributed memory machine, Algorithmica 16 (1996), no. 4-5, 517–542. MR 1407587

## Bibliography IV

- K. Kenthapadi and R. Panigrahy, *Balanced allocation on graphs*, 17th Annual ACM-SIAM Symposium on Discrete Algorithms (SODA'06), ACM, 2006, pp. 434–443. MR 2368840
- D. Los and T. Sauerwald, Balanced allocations in batches: Simplified and generalized, 34th Annual ACM Symposium on Parallel Algorithms and Architectures (SPAA'22), ACM, 2022, p. 389–399.
- Y. Lu, Q. Xie, G. Kliot, A. Geller, J. R. Larus, and A. G. Greenberg, *Join-idle-queue:* A novel load balancing algorithm for dynamically scalable web services, Perform. Evaluation 68 (2011), no. 11, 1056–1071.
- ▶ M. Mitzenmacher, *The power of two choices in randomized load balancing*, Ph.D. thesis, University of California at Berkeley, 1996.
- How useful is old information?, IEEE Trans. Parallel Distributed Syst. 11 (2000), no. 1, 6–20.

## Bibliography V

- ▶ M. Mitzenmacher, B. Prabhakar, and D. Shah, *Load balancing with memory*, The 43rd Annual IEEE Symposium on Foundations of Computer Science, 2002. Proceedings., IEEE, 2002, pp. 799–808.
- ▶ G. Nadiradze, On achieving scalability through relaxation, Ph.D. thesis, IST Austria, 2021.
- K. Ousterhout, P. Wendell, M. Zaharia, and I. Stoica, Sparrow: distributed, low latency scheduling, 24th ACM SIGOPS Symposium on Operating Systems Principles (SOSP'13), ACM, 2013, pp. 69–84.
- ▶ Y. Peres, K. Talwar, and U. Wieder, Graphical balanced allocations and the  $(1 + \beta)$ -choice process, Random Structures & Algorithms **47** (2015), no. 4, 760–775. MR 3418914
- M. Raab and A. Steger, "Balls into bins"—a simple and tight analysis, 2nd International Workshop on Randomization and Computation (RANDOM'98), vol. 1518, Springer, 1998, pp. 159–170. MR 1729169

## Bibliography VI

- J. Spencer, *Balancing games*, J. Combinatorial Theory Ser. B 23 (1977), no. 1, 68–74. MR 526057
- ▶ W. Whitt, *Deciding which queue to join: Some counterexamples*, Oper. Res. **34** (1986), no. 1, 55–62.

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