#### **Balanced Allocations under Incomplete Information**

<u>Dimitrios Los</u><sup>1</sup>, Thomas Sauerwald<sup>1</sup>, John Sylvester<sup>2</sup>

<sup>1</sup>University of Cambridge, UK <sup>2</sup>University of Glasgow, UK

### Balanced allocations: Background

Allocate m tasks (balls) sequentially into n machines (bins).

Allocate m tasks (balls) sequentially into n machines (bins).

**<u>Goal</u>**: minimise the maximum load  $\max_{i \in [n]} x_i^m$ , where  $x^t$  is the load vector after ball t.

Allocate m tasks (balls) sequentially into n machines (bins).

**<u>Goal</u>**: minimise the maximum load  $\max_{i \in [n]} x_i^m$ , where  $x^t$  is the load vector after ball t.



Allocate m tasks (balls) sequentially into n machines (bins).

**<u>Goal</u>**: minimise the maximum load  $\max_{i \in [n]} x_i^m$ , where  $x^t$  is the load vector after ball t.



Allocate m tasks (balls) sequentially into n machines (bins).

**<u>Goal</u>:** minimise the **maximum load**  $\max_{i \in [n]} x_i^m$ , where  $x^t$  is the load vector after ball t.  $\Leftrightarrow$  minimise the **gap**, where  $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$ .



Allocate m tasks (balls) sequentially into n machines (bins).

**<u>Goal</u>:** minimise the maximum load  $\max_{i \in [n]} x_i^m$ , where  $x^t$  is the load vector after ball t.  $\Leftrightarrow$  minimise the gap, where  $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$ .



Allocate m tasks (balls) sequentially into n machines (bins).

**<u>Goal</u>:** minimise the maximum load  $\max_{i \in [n]} x_i^m$ , where  $x^t$  is the load vector after ball t.  $\Leftrightarrow$  minimise the gap, where  $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$ .



#### Applications in hashing, load balancing and routing.

### Outline of the presentation

- **Part A:** Definition of ONE-CHOICE, TWO-CHOICE, and the  $(1 + \beta)$  process.
- **Part B:** The QUANTILE Process
- **Part C:** The MEAN-THRESHOLD Process
- **Part D:** Applications: Outdated information and Noise

<u>ONE-CHOICE Process</u>: Iteration: For each  $t \ge 0$ , sample one bin uniformly at random (u.a.r.) and place the ball there.

<u>ONE-CHOICE Process</u>: Iteration: For each  $t \ge 0$ , sample one bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81].

<u>ONE-CHOICE Process</u>: Iteration: For each  $t \ge 0$ , sample one bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case 
$$(m = n)$$
, w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81].  
Meaning with probability  
at least  $1 - n^{-c}$  for constant  $c > 0$ .

<u>ONE-CHOICE Process</u>: **Iteration**: For each  $t \ge 0$ , sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81].

In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$  (e.g. [RS98]).

<u>ONE-CHOICE Process</u>: **Iteration**: For each  $t \ge 0$ , sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81]. In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$  (e.g. [RS98]).

<u>Two-Choice Process</u>: **Iteration**: For each  $t \ge 0$ , sample **two** bins independently u.a.r. and place the ball in the least loaded of the two.

<u>ONE-CHOICE Process</u>: **Iteration**: For each  $t \ge 0$ , sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case 
$$(m = n)$$
, w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81].  
In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$  (e.g. [RS98]).

<u>TWO-CHOICE Process</u>: Iteration: For each  $t \ge 0$ , sample two bins independently u.a.r. and place the ball in the least loaded of the two.

In the lightly-loaded case (m = n), w.h.p.  $Gap(n) = log_2 log n + \Theta(1)$  [KLMadH96, ABKU99].

<u>ONE-CHOICE Process</u>: **Iteration**: For each  $t \ge 0$ , sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81]. In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n}} \cdot \log n\right)$  (e.g. [RS98]).

<u>Two-Choice Process</u>: **Iteration**: For each  $t \ge 0$ , sample **two** bins independently u.a.r. and place the ball in the least loaded of the two.

In the lightly-loaded case (m = n), w.h.p.  $Gap(n) = log_2 log n + \Theta(1)$  [KLMadH96, ABKU99].

<u>ONE-CHOICE Process</u>: **Iteration**: For each  $t \ge 0$ , sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case 
$$(m = n)$$
, w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81].  
In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$  (e.g. [RS98]).

TWO-CHOICE Process:

**Iteration**: For each  $t \ge 0$ , sample **two** bins independently u.a.r. and place the ball in the least loaded of the two.

- In the lightly-loaded case (m = n), w.h.p.  $Gap(n) = log_2 log n + \Theta(1)$ [KLMadH96, ABKU99].
- In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \log_2 \log n + \Theta(1)$  [BCSV06].

Balanced allocations: Background

<u>ONE-CHOICE Process</u>: **Iteration**: For each  $t \ge 0$ , sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p.  $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$  [Gon81]. In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$  (e.g. [RS98]).

<u>Two-Choice Process</u>: **Iteration**: For each  $t \ge 0$ , sample two bins independently u.a.r. and place the ball in the least loaded of the two.

In the lightly-loaded case (m = n), w.h.p.  $\operatorname{Gap}(n) = \log_2 \log n + \bigoplus(1)$  [KLMadH96, ABKU99].

In the heavily-loaded case  $(m \gg n)$ , w.h.p.  $\operatorname{Gap}(m) = \log_2 \log n + \Theta(1)$  [BCSV06].

# $(1+\beta)$ process: Definition

 $(1 + \beta)$  process: **Parameter**: A probability  $\beta \in (0, 1]$ . **Iteration**: For each  $t \ge 0$ , with probability  $\beta$  allocate one ball via the TWO-CHOICE process, otherwise allocate one ball via the ONE-CHOICE process.

# $(1+\beta)$ process: Definition

 $(1 + \beta)$  process: Parameter: A probability  $\beta \in (0, 1]$ . Iteration: For each  $t \ge 0$ , with probability  $\beta$  allocate one ball via the TWO-CHOICE process, otherwise allocate one ball via the ONE-CHOICE process.

[Mit99] interpreted  $(1 - \beta)/2$  as the probability of making an erroneous comparison.

# $(1+\beta)$ process: Definition

 $(1 + \beta)$  process: Parameter: A probability  $\beta \in (0, 1]$ . Iteration: For each  $t \ge 0$ , with probability  $\beta$  allocate one ball via the Two-CHOICE process, otherwise allocate one ball via the ONE-CHOICE process.

[Mit99] interpreted  $(1 - \beta)/2$  as the probability of making an erroneous comparison.

In the heavily-loaded case, [PTW15] proved that the gap is w.h.p.  $\Theta(\log n/\beta)$  for  $\beta < 1 - \epsilon$  for constant  $\epsilon > 0$ .

### $\mathbf{The} \ \mathbf{Q} \mathbf{UANTILE} \ \mathbf{Process}$

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } \operatorname{Rank}(x^t, i_1) > \delta \cdot n, \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{otherwise.} \end{cases}$$

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } \operatorname{Rank}(x^t, i_1) > \delta \cdot n, \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{otherwise.} \end{cases}$$



$$\begin{aligned} x_{i_1}^{t+1} &= x_{i_1}^t + 1 & \text{if } \operatorname{Rank}(x^t, i_1) > \delta \cdot n, \\ x_{i_2}^{t+1} &= x_{i_2}^t + 1 & \text{otherwise.} \end{aligned}$$



$$\begin{aligned} x_{i_1}^{t+1} &= x_{i_1}^t + 1 & \text{if } \operatorname{Rank}(x^t, i_1) > \delta \cdot n \\ x_{i_2}^{t+1} &= x_{i_2}^t + 1 & \text{otherwise.} \end{aligned}$$



$$\begin{aligned} x_{i_1}^{t+1} &= x_{i_1}^t + 1 & \text{if } \operatorname{Rank}(x^t, i_1) > \delta \cdot n, \\ x_{i_2}^{t+1} &= x_{i_2}^t + 1 & \text{otherwise.} \end{aligned}$$



$$\begin{aligned} x_{i_1}^{t+1} &= x_{i_1}^t + 1 & \text{if } \operatorname{Rank}(x^t, i_1) > \delta \cdot n, \\ x_{i_2}^{t+1} &= x_{i_2}^t + 1 & \text{otherwise.} \end{aligned}$$



$$\begin{aligned} x_{i_1}^{t+1} &= x_{i_1}^t + 1 & \text{if } \operatorname{Rank}(x^t, i_1) > \delta \cdot n, \\ x_{i_2}^{t+1} &= x_{i_2}^t + 1 & \text{otherwise.} \end{aligned}$$



# Quantile( $\delta$ ) as Two-Choice with incomplete information

We can interpret  $QUANTILE(\delta)$  as an instance of the TWO-CHOICE process, where we are only able to compare the loads of the two sampled bins if one is above the quantile and one is below.



# Quantile( $\delta$ ) as Two-Choice with incomplete information

We can interpret  $QUANTILE(\delta)$  as an instance of the TWO-CHOICE process, where we are only able to compare the loads of the two sampled bins if one is above the quantile and one is below.



# Quantile( $\delta$ ) as Two-Choice with incomplete information

We can interpret  $QUANTILE(\delta)$  as an instance of the TWO-CHOICE process, where we are only able to compare the loads of the two sampled bins if one is above the quantile and one is below.



### Probability allocation vectors

**Probability allocation vector**  $p^t$ , where  $p_i^t$  is the prob. of allocating to *i*-th most loaded bin.

#### Probability allocation vectors

- **Probability allocation vector**  $p^t$ , where  $p_i^t$  is the prob. of allocating to *i*-th most loaded bin.
- For ONE-CHOICE,  $p_{\text{ONE-CHOICE}} = \left(\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n}\right)$ .



#### Probability allocation vectors

- **Probability allocation vector**  $p^t$ , where  $p_i^t$  is the prob. of allocating to *i*-th most loaded bin.
- For ONE-CHOICE,  $p_{\text{ONE-CHOICE}} = \left(\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n}\right)$ .
- For Two-CHOICE,  $p_{\text{TWO-CHOICE}} = \left(\frac{1}{n^2}, \frac{3}{n^2}, \dots, \frac{2i-1}{n^2}, \dots, \frac{2n-2}{n^2}\right).$


# Probability allocation vectors

**Probability allocation vector**  $p^t$ , where  $p_i^t$  is the prob. of allocating to *i*-th most loaded bin.



■ [PTW15] used the two-sided **exponential potential** 



■ [PTW15] used the two-sided **exponential potential** 

$$\Gamma^{t}(x^{t}) := \underbrace{\sum_{i=1}^{n} e^{\alpha(x_{i}^{t} - t/n)}}_{\text{Overload potential}} + \underbrace{\sum_{i=1}^{n} e^{-\alpha(x_{i}^{t} - t/n)}}_{\text{Underload potential}}.$$
[PTW15] show that  $\mathbf{E} \left[ \Gamma^{t+1} \mid \mathfrak{F}^{t} \right] \leq \Gamma^{t} \cdot \left(1 - \frac{c_{1}}{n}\right) + c_{2}.$ 

■ [PTW15] used the two-sided **exponential potential** 

$$\Gamma^{t}(x^{t}) := \underbrace{\sum_{i=1}^{n} e^{\alpha(x_{i}^{t} - t/n)}}_{\text{Overload potential}} + \underbrace{\sum_{i=1}^{n} e^{-\alpha(x_{i}^{t} - t/n)}}_{\text{Underload potential}}$$

$$[\text{PTW15] show that \mathbf{E} \left[ \Gamma^{t+1} \mid \mathfrak{F}^{t} \right] \leq \Gamma^{t} \cdot \left( 1 - \frac{c_{1}}{n} \right) + c_{2}.$$

$$\text{This implies } \mathbf{E} \left[ \Gamma^{t} \right] \leq c \cdot n \text{ for any } t \geq 0.$$

.

■ [PTW15] used the two-sided **exponential potential** 

$$\Gamma^{t}(x^{t}) := \sum_{\substack{i=1 \\ \text{Overload potential}}}^{n} e^{\alpha(x_{i}^{t}-t/n)} + \sum_{\substack{i=1 \\ \text{Underload potential}}}^{n} e^{-\alpha(x_{i}^{t}-t/n)} .$$

$$[\text{PTW15] show that } \mathbf{E} \left[ \Gamma^{t+1} \mid \mathfrak{F}^{t} \right] \leq \Gamma^{t} \cdot \left( 1 - \frac{c_{1}}{n} \right) + c_{2}.$$

$$\text{This implies } \mathbf{E} \left[ \Gamma^{t} \right] \leq c \cdot n \text{ for any } t \geq 0.$$

$$\text{By Markov's inequality, we get } \mathbf{Pr} \left[ \Gamma^{m} \leq cn^{3} \right] \geq 1 - n^{-2} \text{ which implies}$$

$$\mathbf{Pr} \left[ \text{Gap}(m) \leq \frac{1}{\alpha} (3 \cdot \log n + \log c) \right] \geq 1 - n^{-2}.$$

■ [PTW15] used the two-sided **exponential potential** 

$$\Gamma^{t}(x^{t}) := \sum_{\substack{i=1 \\ \text{Overload potential}}}^{n} e^{\alpha(x_{i}^{t}-t/n)} + \sum_{\substack{i=1 \\ \text{Underload potential}}}^{n} e^{-\alpha(x_{i}^{t}-t/n)} .$$

$$[\text{PTW15] show that } \mathbf{E} \left[ \Gamma^{t+1} \mid \mathfrak{F}^{t} \right] \leq \Gamma^{t} \cdot \left( 1 - \frac{c_{1}}{n} \right) + c_{2}.$$

$$[\text{This implies } \mathbf{E} \left[ \Gamma^{t} \right] \leq c \cdot n \text{ for any } t \geq 0.$$

$$[\text{By Markov's inequality, we get } \mathbf{Pr} \left[ \Gamma^{m} \leq cn^{3} \right] \geq 1 - n^{-2} \text{ which implies}$$

$$[\text{Pr} \left[ \text{Gap}(m) \leq \frac{1}{\alpha} (3 \cdot \log n + \log c) \right] \geq 1 - n^{-2}.$$

For the  $(1 + \beta)$  process,  $\alpha = \Theta(\beta)$ .

■ [PTW15] used the two-sided **exponential potential** 

$$\Gamma^{t}(x^{t}) := \sum_{\substack{i=1\\\text{Overload potential}}}^{n} e^{\alpha(x_{i}^{t}-t/n)} + \sum_{\substack{i=1\\\text{Underload potential}}}^{n} e^{-\alpha(x_{i}^{t}-t/n)} .$$

$$[\text{PTW15] show that } \mathbf{E} \left[ \Gamma^{t+1} \mid \mathfrak{F}^{t} \right] \leq \Gamma^{t} \cdot \left( 1 - \frac{c_{1}}{n} \right) + c_{2}.$$

$$[\text{This implies } \mathbf{E} \left[ \Gamma^{t} \right] \leq c \cdot n \text{ for any } t \geq 0.$$

$$[\text{By Markov's inequality, we get } \mathbf{Pr} \left[ \Gamma^{m} \leq cn^{3} \right] \geq 1 - n^{-2} \text{ which implies}$$

$$[\text{Pr} \left[ \text{Gap}(m) \leq \frac{1}{\alpha} (3 \cdot \log n + \log c) \right] \geq 1 - n^{-2}.$$

For the (1 + β) process, α = Θ(β).
Same proof holds for the QUANTILE(δ) for constant δ ∈ (0, 1).

■ [PTW15] used the two-sided **exponential potential** 

$$\Gamma^{t}(x^{t}) := \sum_{\substack{i=1 \\ \text{Overload potential}}}^{n} e^{\alpha(x_{i}^{t}-t/n)} + \sum_{\substack{i=1 \\ \text{Underload potential}}}^{n} e^{-\alpha(x_{i}^{t}-t/n)} .$$

$$[\text{PTW15] show that } \mathbf{E} \left[ \Gamma^{t+1} \mid \mathfrak{F}^{t} \right] \leq \Gamma^{t} \cdot \left( 1 - \frac{c_{1}}{n} \right) + c_{2}.$$

$$[\text{This implies } \mathbf{E} \left[ \Gamma^{t} \right] \leq c \cdot n \text{ for any } t \geq 0.$$

$$[\text{By Markov's inequality, we get } \mathbf{Pr} \left[ \Gamma^{m} \leq cn^{3} \right] \geq 1 - n^{-2} \text{ which implies}$$

$$[\text{Pr} \left[ \text{Gap}(m) \leq \frac{1}{\alpha} (3 \cdot \log n + \log c) \right] \geq 1 - n^{-2}.$$

For the  $(1 + \beta)$  process,  $\alpha = \Theta(\beta)$ .

Same proof holds for the QUANTILE( $\delta$ ) for constant  $\delta \in (0, 1)$ .

In [PTW15],  $\alpha = \mathcal{O}(1)$  so the tightest gaps proved were  $\mathcal{O}(\log n)$ .

# QUANTILE $(\delta_1, \ldots, \delta_k)$ process

We can extend the QUANTILE( $\delta$ ) process to k quantiles.

# Quantile $(\delta_1, \ldots, \delta_k)$ process

We can extend the  $QUANTILE(\delta)$  process to k quantiles.

We can only distinguish two bins if they are in different regions.



# Quantile $(\delta_1, \ldots, \delta_k)$ process

We can extend the  $QUANTILE(\delta)$  process to k quantiles.

We can only distinguish two bins if they are in different regions.



# Quantile $(\delta_1, \ldots, \delta_k)$ process

We can extend the  $QUANTILE(\delta)$  process to k quantiles.

We can only distinguish two bins if they are in different regions.



A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .

A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .



A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .



A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .



A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .



A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .



A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .

Implications:

A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .

Implications:

► For  $k = \Theta(\log \log n)$ , we recover the **Two-CHOICE** Gap $(m) = \mathcal{O}(\log \log n)$ .

A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .

Implications:

► For  $k = \Theta(\log \log n)$ , we recover the Two-CHOICE  $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$ .

For 
$$(1 + \beta)$$
 with  $\beta = 1 - 2^{-0.5(\log n)^{(k-1)/k}}$ , w.h.p.  $\operatorname{Gap}(m) = \mathcal{O}(k \cdot (\log n)^{1/k})$ .

A QUANTILE $(\delta_1, \ldots, \delta_k)$  process with uniform quantiles that achieves w.h.p. an  $\mathcal{O}(k \cdot (\log n)^{1/k})$  gap for  $k = \mathcal{O}(\log \log n)$ .

Implications:

- ► For  $k = \Theta(\log \log n)$ , we recover the Two-CHOICE  $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$ .
- For  $(1 + \beta)$  with  $\beta = 1 2^{-0.5(\log n)^{(k-1)/k}}$ , w.h.p.  $\operatorname{Gap}(m) = \mathcal{O}(k \cdot (\log n)^{1/k})$ .
- ▶ Improvements for other processes (*d*-THINNING, graphical allocations).









For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

$$\Phi_0^t := \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right),$$

For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

$$\begin{split} \Phi_0^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right), \\ \Phi_1^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{1/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{2}{\alpha'} (\log n)^{1/3}\right)^+\right), \end{split}$$

For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

$$\begin{split} \Phi_0^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right), \\ \Phi_1^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{1/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{2}{\alpha'} (\log n)^{1/3}\right)^+\right), \\ \Phi_2^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{2/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{4}{\alpha'} (\log n)^{1/3}\right)^+\right). \end{split}$$

For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

$$\begin{split} \Phi_0^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right), & \mathcal{O}(\log n) \\ \Phi_1^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{1/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{2}{\alpha'} (\log n)^{1/3}\right)^+\right), \\ \Phi_2^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{2/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{4}{\alpha'} (\log n)^{1/3}\right)^+\right). \end{split}$$

For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

$$\begin{split} \Phi_0^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right), & \mathcal{O}(\log n) \\ \Phi_1^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{1/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{2}{\alpha'} (\log n)^{1/3}\right)^+\right), & \mathcal{O}((\log n)^{2/3}) \\ \Phi_2^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{2/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{4}{\alpha'} (\log n)^{1/3}\right)^+\right). \end{split}$$

For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

$$\begin{split} \Phi_0^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right), & \mathcal{O}(\log n) \\ \Phi_1^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{1/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{2}{\alpha'} (\log n)^{1/3}\right)^+\right), & \mathcal{O}((\log n)^{2/3}) \\ \Phi_2^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{2/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{4}{\alpha'} (\log n)^{1/3}\right)^+\right). & \mathcal{O}((\log n)^{1/3}) \end{split}$$

For k = 3, we define the quantile process which achives  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

• We define the following exponential potential functions:

$$\begin{split} \Phi_0^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right), & \mathcal{O}(\log n) \\ \Phi_1^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{1/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{2}{\alpha'} (\log n)^{1/3}\right)^+\right), & \mathcal{O}((\log n)^{2/3}) \\ \Phi_2^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{2/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{4}{\alpha'} (\log n)^{1/3}\right)^+\right). & \mathcal{O}((\log n)^{1/3}) \end{split}$$

When  $\alpha = \Omega(1)$ , the potential may not necessarily drop in expectation.

For k = 3, we define the quantile process which achieves  $\mathcal{O}((\log n)^{1/3})$  gap:  $\delta_1 := e^{-\Theta((\log n)^{1/3})}, \quad \delta_2 := e^{-\Theta((\log n)^{2/3})}, \quad \delta_3 := 1/2$ 

• We define the following exponential potential functions:

$$\begin{split} \Phi_0^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot \left(x_i^t - \frac{t}{n}\right)^+\right), & \mathcal{O}(\log n) \\ \Phi_1^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{1/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{2}{\alpha'}(\log n)^{1/3}\right)^+\right), & \mathcal{O}((\log n)^{2/3}) \\ \Phi_2^t &:= \sum_{i=1}^n \exp\left(\alpha' \cdot (\log n)^{2/3} \cdot \left(x_i^t - \frac{t}{n} - \frac{4}{\alpha'}(\log n)^{1/3}\right)^+\right). & \mathcal{O}((\log n)^{1/3}) \\ \text{When } \alpha &= \Omega(1), \text{ the potential may not necessarily drop in expectation.} \\ \text{We prove that when } y_{\delta_{3-j}\cdot n}^t < \frac{2}{\alpha'}j(\log n)^{1/3}, \text{ then} \\ & \mathbf{E}\left[\Phi_j^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi_j^t \cdot \left(1 - \frac{1}{n}\right) + 2. \end{split}$$

**Proving**  $\operatorname{Gap}(m) = \mathcal{O}(k \cdot (\log n)^{1/k})$ 














# Mean-Threshold

THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$

THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$



THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$

THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$



THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$



THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$



THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$

• MEAN-THRESHOLD has f(n) = 0.

[FGG21] found the asymptotically optimal threshold in the lightly-loaded case.

THRESHOLD(f(n)) Process:

**Parameter**: A threshold function  $f(n) \ge 0$ .

**Iteration**: For  $t \ge 0$ , sample two uniform bins  $i_1$  and  $i_2$  independently, and update:

$$\begin{cases} x_{i_1}^{t+1} = x_{i_1}^t + 1 & \text{if } x_{i_1}^t < \frac{t}{n} + f(n), \\ x_{i_2}^{t+1} = x_{i_2}^t + 1 & \text{if } x_{i_1}^t \ge \frac{t}{n} + f(n). \end{cases}$$

MEAN-THRESHOLD has f(n) = 0.

[FGG21] found the asymptotically optimal threshold in the lightly-loaded case.

**[IK05, FL20]** analysed a *d*-sample version for the lightly-loaded case.

### THRESHOLD as TWO-CHOICE with incomplete information

We can interpret **THRESHOLD** as an instance of the TWO-CHOICE process, where we are only able to compare the loads of the two sampled bins if one is above the threshold and one is below.



### THRESHOLD as TWO-CHOICE with incomplete information

We can interpret **THRESHOLD** as an instance of the TWO-CHOICE process, where we are only able to compare the loads of the two sampled bins if one is above the threshold and one is below.



### THRESHOLD as TWO-CHOICE with incomplete information

We can interpret **THRESHOLD** as an instance of the TWO-CHOICE process, where we are only able to compare the loads of the two sampled bins if one is above the threshold and one is below.



**MEAN-THRESHOLD** does not require communication of the load.

**MEAN-THRESHOLD** does not require communication of the load.

**MEAN-THRESHOLD** does not require communication of the load.



**MEAN-THRESHOLD** does not require communication of the load.



**MEAN-THRESHOLD** does not require communication of the load.



**MEAN-THRESHOLD** does not require communication of the load.



- **MEAN-THRESHOLD** does not require communication of the load.
- Requires 1-bit responses.
- Or we can completely avoid responses to the allocator.



- **MEAN-THRESHOLD** does not require communication of the load.
- Requires 1-bit responses.
- Or we can completely avoid responses to the allocator.



- **MEAN-THRESHOLD** does not require communication of the load.
- Requires 1-bit responses.
- Or we can completely avoid responses to the allocator.



### MEAN-THRESHOLD: Our results

For heavily-loaded case, MEAN-THRESHOLD achieves w.h.p.  $Gap(m) = \mathcal{O}(\log n)$ .

### MEAN-THRESHOLD: Our results

For heavily-loaded case, MEAN-THRESHOLD achieves w.h.p.  $\operatorname{Gap}(m) = \mathcal{O}(\log n)$ .

For sufficiently large m, MEAN-THRESHOLD achieves w.h.p.  $Gap(m) = \Omega(\log n)$ .

### MEAN-THRESHOLD: Our results

- For heavily-loaded case, MEAN-THRESHOLD achieves w.h.p.  $\operatorname{Gap}(m) = \mathcal{O}(\log n)$ .
- For sufficiently large m, MEAN-THRESHOLD achieves w.h.p.  $Gap(m) = \Omega(\log n)$ .
- **MEAN-THRESHOLD** uses w.h.p.  $2 \epsilon$  samples per allocation.

In Let  $\delta^t$  be the quantile position of the mean.

In Let  $\delta^t$  be the quantile position of the mean.

If  $\delta^t$  is very large, say  $\delta^t = 1 - 1/n$ , then p becomes very close to the ONE-CHOICE vector :

$$p_{\text{MEAN-THRESHOLD}}(x^{t}) = \left(\underbrace{\frac{1}{n} - \frac{1}{n^{2}}, \dots, \frac{1}{n} - \frac{1}{n^{2}}}_{(n-1) \text{ entries}}, \frac{2}{n} - \frac{1}{n^{2}}\right).$$

 $\blacksquare$  Let  $\delta^t$  be the quantile position of the mean.

If  $\delta^t$  is very large, say  $\delta^t = 1 - 1/n$ , then p becomes very close to the ONE-CHOICE vector :

$$p_{\text{MEAN-THRESHOLD}}(x^{t}) = \left(\underbrace{\frac{1}{n} - \frac{1}{n^{2}}, \dots, \frac{1}{n} - \frac{1}{n^{2}}}_{(n-1) \text{ entries}}, \frac{2}{n} - \frac{1}{n^{2}}\right).$$

With this worst-case probability vector, we can only obtain w.h.p. a gap of  $\mathcal{O}(n \log n)$ using  $\Gamma^t$  with  $\alpha = \Theta(1/n)$ .

 $\blacksquare$  Let  $\delta^t$  be the quantile position of the mean.

If  $\delta^t$  is very large, say  $\delta^t = 1 - 1/n$ , then p becomes very close to the ONE-CHOICE vector :

$$p_{\text{MEAN-THRESHOLD}}(x^{t}) = \left(\underbrace{\frac{1}{n} - \frac{1}{n^{2}}, \dots, \frac{1}{n} - \frac{1}{n^{2}}}_{(n-1) \text{ entries}}, \frac{2}{n} - \frac{1}{n^{2}}\right).$$

With this worst-case probability vector, we can only obtain w.h.p. a gap of  $\mathcal{O}(n \log n)$ using  $\Gamma^t$  with  $\alpha = \Theta(1/n)$ .

But what happens for  $\Gamma^t$  with constant  $\alpha$ ?

# ${\it Mean-Threshold:} {\bf Bad \ configuration}$



There is a very small bias away from overloaded bins.

The exponential potential for constant  $\alpha$  increases in expectation.

### MEAN-THRESHOLD: Recovery from a bad configuration
An analysis similar to [PTW15] shows that

An analysis similar to [PTW15] shows that  $(\mathbf{C} = \mathbf{1} + \mathbf{1}) = \mathbf{1} + \mathbf{1}$ 

▶ (Good step) If  $\delta^t \in (\epsilon, 1 - \epsilon)$  for const  $\epsilon > 0$ , then

$$\mathbf{E}[\,\Gamma^{t+1} \mid \mathfrak{F}^t, \{\delta^t \in (\epsilon, 1-\epsilon)\}, \Gamma^t \ge c \cdot n\,] \le \Gamma^t \cdot \left(1 - \Theta\!\left(\frac{\alpha}{n}\right)\right).$$

An analysis similar to [PTW15] shows that
 (Good step) If δ<sup>t</sup> ∈ (ε, 1 − ε) for const ε > 0, then

$$\mathbf{E}[\,\Gamma^{t+1} \mid \mathfrak{F}^t, \{\delta^t \in (\epsilon, 1-\epsilon)\}, \Gamma^t \ge c \cdot n\,] \le \Gamma^t \cdot \left(1 - \Theta\!\left(\frac{\alpha}{n}\right)\right).$$

▶ (**Bad step**) If 
$$\delta^t \notin (\epsilon, 1 - \epsilon)$$
, then

$$\mathbf{E}[\Gamma^{t+1} \mid \mathfrak{F}^t, \Gamma^t \ge c \cdot n] \le \Gamma^t \cdot \left(1 + \Theta\left(\frac{\alpha^2}{n}\right)\right).$$

An analysis similar to [PTW15] shows that

▶ (Good step) If  $\delta^t \in (\epsilon, 1 - \epsilon)$  for const  $\epsilon > 0$ , then

$$\mathbf{E}[\,\Gamma^{t+1} \mid \mathfrak{F}^t, \{\delta^t \in (\epsilon, 1-\epsilon)\}, \Gamma^t \ge c \cdot n\,] \le \Gamma^t \cdot \left(1 - \Theta\!\left(\frac{\alpha}{n}\right)\right).$$

▶ (**Bad step**) If 
$$\delta^t \notin (\epsilon, 1 - \epsilon)$$
, then

$$\mathbf{E}[\Gamma^{t+1} \mid \mathfrak{F}^t, \Gamma^t \ge c \cdot n] \le \Gamma^t \cdot \left(1 + \Theta\left(\frac{\alpha^2}{n}\right)\right).$$

A properly tweaked potential function drops in expectation for any interval with constant fraction of good steps.

An analysis similar to [PTW15] shows that

▶ (Good step) If  $\delta^t \in (\epsilon, 1 - \epsilon)$  for const  $\epsilon > 0$ , then

$$\mathbf{E}[\,\Gamma^{t+1} \mid \mathfrak{F}^t, \{\delta^t \in (\epsilon, 1-\epsilon)\}, \Gamma^t \ge c \cdot n\,] \le \Gamma^t \cdot \left(1 - \Theta\!\left(\frac{\alpha}{n}\right)\right).$$

▶ (**Bad step**) If 
$$\delta^t \notin (\epsilon, 1 - \epsilon)$$
, then

$$\mathbf{E}[\Gamma^{t+1} \mid \mathfrak{F}^t, \Gamma^t \ge c \cdot n] \le \Gamma^t \cdot \left(1 + \Theta\left(\frac{\alpha^2}{n}\right)\right).$$

• A properly tweaked potential function drops in expectation for any interval with constant fraction of good steps.

How can we prove that there is a constant fraction of good steps?

Consider the **absolute value** and **quadratic** potentials,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right| \quad \text{and} \quad \Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$



Consider the **absolute value** and **quadratic** potentials,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right| \quad \text{and} \quad \Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$



As long as  $\Delta^t = \Omega(n)$ ,  $\Upsilon^t$  drops in expectation.

Consider the **absolute value** and **quadratic** potentials,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right| \quad \text{and} \quad \Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$



Mean-Threshold

Consider the **absolute value** and **quadratic** potentials,



As  $\Delta^t$  becomes smaller,  $\delta^t$  improves and  $\Gamma^t$  drops in expectation. MEAN-THRESHOLD

Summary of results:

Analysed various processes including *k*-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.

- Analysed various processes including *k*-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:

- Analysed various processes including *k*-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.

- Analysed various processes including k-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.
  - $\blacktriangleright$  Interplay between the absolute value and quadratic potentials.

- Analysed various processes including k-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.
  - ▶ Interplay between the absolute value and quadratic potentials.
- Further applications of the presented techniques:
  - ► Balls allocated in batches of  $b \ge n$  balls :  $\mathcal{O}(b/n + \log n)$  gap for Two-CHOICE,  $(1 + \beta)$  and QUANTILE( $\delta$ ) [LS22a].

- Analysed various processes including k-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.
  - ▶ Interplay between the absolute value and quadratic potentials.
- Further applications of the presented techniques:
  - ▶ Balls allocated in batches of  $b \ge n$  balls :  $\mathcal{O}(b/n + \log n)$  gap for Two-CHOICE,  $(1 + \beta)$  and QUANTILE( $\delta$ ) [LS22a].
  - ▶ Two-CHOICE with adversarial noise  $g \leq \log n$  :  $\mathcal{O}(\frac{g}{\log q} \cdot \log \log n)$  [LS22b].

Summary of results:

- Analysed various processes including *k*-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.
  - ▶ Interplay between the absolute value and quadratic potentials.
- Further applications of the presented techniques:
  - ▶ Balls allocated in batches of  $b \ge n$  balls :  $\mathcal{O}(b/n + \log n)$  gap for Two-CHOICE,  $(1 + \beta)$  and QUANTILE( $\delta$ ) [LS22a].
  - ▶ Two-CHOICE with adversarial noise  $g \leq \log n$  :  $\mathcal{O}(\frac{g}{\log q} \cdot \log \log n)$  [LS22b].

Several directions for future work:

Summary of results:

- Analysed various processes including *k*-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.
  - ▶ Interplay between the absolute value and quadratic potentials.
- Further applications of the presented techniques:
  - ► Balls allocated in batches of  $b \ge n$  balls :  $\mathcal{O}(b/n + \log n)$  gap for Two-CHOICE,  $(1 + \beta)$  and QUANTILE( $\delta$ ) [LS22a].
  - ▶ Two-CHOICE with adversarial noise  $g \leq \log n$  :  $\mathcal{O}(\frac{g}{\log q} \cdot \log \log n)$  [LS22b].

Several directions for future work:

Investigating the robustness of MEAN-THRESHOLD.

Summary of results:

- Analysed various processes including *k*-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.
  - ▶ Interplay between the absolute value and quadratic potentials.
- Further applications of the presented techniques:
  - ► Balls allocated in batches of  $b \ge n$  balls :  $\mathcal{O}(b/n + \log n)$  gap for Two-CHOICE,  $(1 + \beta)$  and QUANTILE( $\delta$ ) [LS22a].
  - ▶ Two-CHOICE with adversarial noise  $g \leq \log n$  :  $\mathcal{O}(\frac{g}{\log g} \cdot \log \log n)$  [LS22b].

Several directions for future work:

- Investigating the robustness of MEAN-THRESHOLD.
- Loosening the fully random hash function assumption.

Summary of results:

- Analysed various processes including *k*-QUANTILE, MEAN-THRESHOLD and  $(1 + \beta)$  for  $\beta$  close to 1.
- Introduced two techniques for analysing balanced allocation processes:
  - ▶ Layered induction over super-exponential potentials.
  - ▶ Interplay between the absolute value and quadratic potentials.
- Further applications of the presented techniques:
  - ► Balls allocated in batches of  $b \ge n$  balls :  $\mathcal{O}(b/n + \log n)$  gap for Two-CHOICE,  $(1 + \beta)$  and QUANTILE( $\delta$ ) [LS22a].
  - ▶ Two-CHOICE with adversarial noise  $g \leq \log n$  :  $\mathcal{O}(\frac{g}{\log g} \cdot \log \log n)$  [LS22b].

Several directions for future work:

- Investigating the robustness of MEAN-THRESHOLD.
- Loosening the fully random hash function assumption.
- Analysing these processes in the graphical setting.

#### **Questions?**

More visualisations: tinyurl.com/lss21-visualisations

#### **Questions?**

More visualisations: tinyurl.com/lss21-visualisations

# Appendix

# Appendix A: Table of results

Process	Lightly Loaded Case $m = \mathcal{O}(n)$			Heavily Loaded Case $m = \omega(n)$		
1 100055	Lower Bound	ower Bound Upper Bound		Lower Bound		Upper Bound
$(1+\beta)$ , const $\beta \in (0,1)$	$\frac{\log n}{\log \log n}$ [PTW15]		$\log n$			
Caching	$\log \log n$		[MPS02]	_		$\log n$
Packing	$\frac{\log n}{\log \log n}$			$\log n$		
TWINNING	$\frac{\log n}{\log \log n}$			$\log n$		
Mean-Threshold	$\frac{\log n}{\log \log n}$			$\log n$		
2-THINNING $\left(\Theta(\sqrt{\frac{\log n}{\log \log n}})\right)$	$\sqrt{10}$	$\frac{\log n}{\log \log n}$	[FL20]	$\frac{\log n}{\log \log n}$	[LS21]	$\log n$
Adaptive-2-Thinning	$\sqrt{10}$	$\frac{\log n}{\log \log n}$	[FL20]	$\frac{\log n}{\log \log n}$	[LS21]	$\frac{\log n}{\log \log n}$ [FGGL21]

**Table:** Overview of the Gap achieved (with probability at least  $1 - n^{-1}$ ), by different allocation processes considered in this work (and related works).

# Appendix B: Detailed experimental results (I)

n	Mean-Threshold	TWINNING	Packing	Caching
$10^{5}$	$\begin{array}{rrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrr$	14 : 2% $15 : 5%$ $16 : 25%$ $17 : 28%$ $18 : 17%$ $19 : 10%$ $20 : 8%$ $21 : 1%$ $22 : 1%$ $23 : 3%$	12 : 2% $13 : 16%$ $14 : 20%$ $15 : 28%$ $16 : 23%$ $17 : 5%$ $18 : 3%$ $19 : 1%$ $20 : 2%$	<b>3</b> : 100%

**Table:** Summary of observed gaps for  $n \in \{10^3, 10^4, 10^5\}$  bins and  $m = 1000 \cdot n$  number of balls, for 100 repetitions. The observed gaps are in bold and next to that is the % of runs where this was observed.

# Appendix B: Detailed experimental results (II)

n	$(1+\beta)$ , for $\beta = 0.5$	k = 1	k = 2	k = 3	k = 4	TWO-CHOICE
$10^{5}$	$\begin{array}{rrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrr$	8:28% 9:42% 10:18% 11:7% 12:3% 14:1% 15:1%	4:72% 5:26% 6:2%	<b>3</b> : 46% <b>4</b> : 54%	$f{3}:79\%\ f{4}:21\%$	<b>3</b> : 100%

**Table:** Summary of our Experimental Results  $(m = 1000 \cdot n)$ .

### Appendix C: Recovery from a bad configuration



We analyze a more general framework that includes PACKING and CACHING [MPS02].

We analyze a more general framework that includes PACKING and CACHING [MPS02].
We prove an O(log n) gap for these processes.

We analyze a more general framework that includes PACKING and CACHING [MPS02].
We prove an O(log n) gap for these processes.



We analyze a more general framework that includes PACKING and CACHING [MPS02].
We prove an O(log n) gap for these processes.











# Appendix F: Outdated information

In "Balanced Allocations in Batches: Simplified and Generalized" [LS22a], we study process with outdated information:

# Appendix F: Outdated information

In "Balanced Allocations in Batches: Simplified and Generalized" [LS22a], we study process with outdated information:

In the *batched setting* balls arrive in batches of size b.
In "Balanced Allocations in Batches: Simplified and Generalized" [LS22a], we study process with outdated information:

In the *batched setting* balls arrive in batches of size b.

For b = n, [BCE<sup>+</sup>12] proved that TWO-CHOICE has w.h.p. an  $\mathcal{O}(\log n)$  gap.

In "Balanced Allocations in Batches: Simplified and Generalized" [LS22a], we study process with outdated information:

- In the *batched setting* balls arrive in batches of size b.
- For b = n, [BCE<sup>+</sup>12] proved that TWO-CHOICE has w.h.p. an  $\mathcal{O}(\log n)$  gap.
- We show that a large class of process (including  $(1 + \beta)$  and QUANTILE( $\delta$ ) for const  $\beta$  and  $\delta$ ) have an  $\mathcal{O}(b/n \cdot \log n)$  gap for any  $b \ge n$ .

In "Balanced Allocations in Batches: Simplified and Generalized" [LS22a], we study process with outdated information:

- In the *batched setting* balls arrive in batches of size *b*.
- For b = n, [BCE<sup>+</sup>12] proved that TWO-CHOICE has w.h.p. an  $\mathcal{O}(\log n)$  gap.
- We show that a large class of process (including  $(1 + \beta)$  and QUANTILE( $\delta$ ) for const  $\beta$  and  $\delta$ ) have an  $\mathcal{O}(b/n \cdot \log n)$  gap for any  $b \ge n$ .
- The proof follows by looking at  $\Gamma$  with  $\alpha = \Theta(n/b)$ .

In "Balanced Allocations in Batches: Simplified and Generalized" [LS22a], we study process with outdated information:

- In the *batched setting* balls arrive in batches of size *b*.
- For b = n, [BCE<sup>+</sup>12] proved that TWO-CHOICE has w.h.p. an  $\mathcal{O}(\log n)$  gap.
- We show that a large class of process (including  $(1 + \beta)$  and QUANTILE( $\delta$ ) for const  $\beta$  and  $\delta$ ) have an  $\mathcal{O}(b/n \cdot \log n)$  gap for any  $b \ge n$ .
- The proof follows by looking at  $\Gamma$  with  $\alpha = \Theta(n/b)$ .
- By using a second potential  $\tilde{\Gamma}$  with  $\tilde{\alpha} = \Theta(\min(1/\log n, n/b))$  and conditioning on  $\Gamma = \mathcal{O}(n)$ , we prove an  $\mathcal{O}(n/b + \log n)$  gap for  $b \ge n$ .

In "Balanced Allocations with the Choice of Noise" [LS22b], we study TWO-CHOICE with noise:

In the *adversarial noise setting*, an adversary can perturb the observed loads by some amount g.

- In the *adversarial noise setting*, an adversary can perturb the observed loads by some amount g.
- Using an interplay between  $\Delta^t$  and  $\Upsilon$ , we prove an  $\mathcal{O}(g + \log n)$  gap.

- In the *adversarial noise setting*, an adversary can perturb the observed loads by some amount g.
- Using an interplay between  $\Delta^t$  and  $\Upsilon$ , we prove an  $\mathcal{O}(g + \log n)$  gap.
- Using layered induction of super-exponential potentials we get  $\mathcal{O}(\frac{g}{\log g} \cdot \log \log n)$  for  $g \leq \log n$ , which is tight.

- In the *adversarial noise setting*, an adversary can perturb the observed loads by some amount g.
- Using an interplay between  $\Delta^t$  and  $\Upsilon$ , we prove an  $\mathcal{O}(g + \log n)$  gap.
- Using layered induction of super-exponential potentials we get  $\mathcal{O}(\frac{g}{\log g} \cdot \log \log n)$  for  $g \leq \log n$ , which is tight.
- Implies tight bounds for random noise from sub-exponential distributions.

- In the *adversarial noise setting*, an adversary can perturb the observed loads by some amount g.
- Using an interplay between  $\Delta^t$  and  $\Upsilon$ , we prove an  $\mathcal{O}(g + \log n)$  gap.
- Using layered induction of super-exponential potentials we get  $\mathcal{O}(\frac{g}{\log g} \cdot \log \log n)$  for  $g \leq \log n$ , which is tight.
- Implies tight bounds for random noise from sub-exponential distributions.
- Implies tight upper bounds for TWO-CHOICE with batch sizes of  $b = \mathcal{O}(n)$ .

- In the *adversarial noise setting*, an adversary can perturb the observed loads by some amount g.
- Using an interplay between  $\Delta^t$  and  $\Upsilon$ , we prove an  $\mathcal{O}(g + \log n)$  gap.
- Using layered induction of super-exponential potentials we get  $\mathcal{O}(\frac{g}{\log g} \cdot \log \log n)$  for  $g \leq \log n$ , which is tight.
- Implies tight bounds for random noise from sub-exponential distributions.
- Implies tight upper bounds for TWO-CHOICE with batch sizes of b = O(n).
- In particular, implies  $\operatorname{Gap}(n) = \Theta(\log n / \log \log n)$  for b = n.
- And for the setting where the load of a bin is chosen adversarially from the last b steps.

Consider the absolute value potential,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right|.$$

Consider the absolute value potential,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right|.$$

If  $\Delta^t = \mathcal{O}(n)$ , then  $\delta^t \in (\epsilon, 1 - \epsilon)$  w.h.p. for a constant fraction of the next  $\Theta(n)$  steps.

Consider the absolute value potential,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right|.$$

If Δ<sup>t</sup> = O(n), then δ<sup>t</sup> ∈ (ε, 1 − ε) w.h.p. for a constant fraction of the next Θ(n) steps.
 Consider the quadratic potential,

$$\Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$

Consider the absolute value potential,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right|.$$

If Δ<sup>t</sup> = O(n), then δ<sup>t</sup> ∈ (ε, 1 − ε) w.h.p. for a constant fraction of the next Θ(n) steps.
 Consider the quadratic potential,

$$\Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$

We prove that

$$\mathbf{E}\left[\,\Upsilon^{t+1}\mid\mathfrak{F}^t\,\right]\leq\Upsilon^t-\Delta^t+1.$$

Consider the absolute value potential,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right|.$$

If Δ<sup>t</sup> = O(n), then δ<sup>t</sup> ∈ (ε, 1 − ε) w.h.p. for a constant fraction of the next Θ(n) steps.
 Consider the quadratic potential,

$$\Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$

We prove that

$$\mathbf{E}\left[\left.\Upsilon^{t+1} \mid \mathfrak{F}^t\right] \leq \Upsilon^t - \Delta^t + 1.$$

By induction we get,

$$\mathbf{E}[\Upsilon^{t+k+1} \mid \mathfrak{F}^t] \leq \Upsilon^t - \frac{1}{n} \cdot \sum_{r=t}^{t+k} \mathbf{E}[\Delta^r \mid \mathfrak{F}^t] + (k+1).$$

Appendix

Consider the absolute value potential,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right|.$$

If Δ<sup>t</sup> = O(n), then δ<sup>t</sup> ∈ (ε, 1 − ε) w.h.p. for a constant fraction of the next Θ(n) steps.
 Consider the quadratic potential,

$$\Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$

We prove that

$$\mathbf{E}\left[\left.\Upsilon^{t+1} \mid \mathfrak{F}^t\right] \leq \Upsilon^t - \Delta^t + 1.$$

By induction we get,

$$\mathbf{E}[\Upsilon^{t+k+1} \mid \mathfrak{F}^t] \leq \Upsilon^t - \frac{1}{n} \cdot \sum_{r=t}^{t+k} \mathbf{E}[\Delta^r \mid \mathfrak{F}^t] + (k+1).$$

Appendix

Consider the absolute value potential,

$$\Delta^t := \sum_{i=1}^n \left| x_i^t - \frac{t}{n} \right|.$$

If Δ<sup>t</sup> = O(n), then δ<sup>t</sup> ∈ (ε, 1 − ε) w.h.p. for a constant fraction of the next Θ(n) steps.
 Consider the quadratic potential,

$$\Upsilon^t := \sum_{i=1}^n \left( x_i^t - \frac{t}{n} \right)^2.$$

We prove that

By induction we get,

$$\begin{split} \mathbf{E}\left[\left.\Upsilon^{t+1} \mid \mathfrak{F}^{t}\right] &\leq \Upsilon^{t} - \Delta^{t} + 1.\\ \text{For } k &= \Theta(\Upsilon^{t}), \text{ for constant fraction of steps } r \in [t, t+k], \ \mathbf{E}\left[\Delta^{r} \mid \mathfrak{F}^{t}\right] = \mathcal{O}(n).\\ \mathbf{E}\left[\Upsilon^{t+k+1} \mid \mathfrak{F}^{t}\right] &\leq \Upsilon^{t} - \frac{1}{n} \cdot \sum_{r=t}^{t+k} \mathbf{E}\left[\Delta^{r} \mid \mathfrak{F}^{t}\right] + (k+1). \end{split}$$

Appendix

# Bibliography I

- Y. Azar, A. Z. Broder, A. R. Karlin, and E. Upfal, *Balanced allocations*, SIAM J. Comput. 29 (1999), no. 1, 180–200. MR 1710347
- Petra Berenbrink, Artur Czumaj, Matthias Englert, Tom Friedetzky, and Lars Nagel, Multiple-choice balanced allocation in (almost) parallel, Proceedings of 16th International Workshop on Approximation, Randomization, and Combinatorial Optimization (RANDOM'12) (Berlin Heidelberg), Springer-Verlag, 2012, pp. 411–422.
- P. Berenbrink, A. Czumaj, A. Steger, and B. Vöcking, Balanced allocations: the heavily loaded case, SIAM J. Comput. 35 (2006), no. 6, 1350–1385. MR 2217150
- Ohad N. Feldheim and Ori Gurel-Gurevich, The power of thinning in balanced allocation, Electron. Commun. Probab. 26 (2021), Paper No. 34, 8. MR 4275960
- O. N. Feldheim, O. Gurel-Gurevich, and J. Li, Long-term balanced allocation via thinning, 2021, arXiv:2110.05009.
- O. N. Feldheim and J. Li, Load balancing under d-thinning, Electron. Commun. Probab.
  25 (2020), Paper No. 1, 13. MR 4053904

# Bibliography II

- ▶ G. H. Gonnet, Expected length of the longest probe sequence in hash code searching, J. Assoc. Comput. Mach. **28** (1981), no. 2, 289–304. MR 612082
- ▶ K. Iwama and A. Kawachi, *Approximated two choices in randomized load balancing*, Algorithms and Computation (Berlin, Heidelberg) (Rudolf Fleischer and Gerhard Trippen, eds.), Springer Berlin Heidelberg, 2005, pp. 545–557.
- R. M. Karp, M. Luby, and F. Meyer auf der Heide, Efficient PRAM simulation on a distributed memory machine, Algorithmica 16 (1996), no. 4-5, 517–542. MR 1407587
- ▶ D. Los and T. Sauerwald, *Balanced allocations with incomplete information: The power* of two queries, 2021, arXiv:2107.03916.
- ▶ Dimitrios Los and Thomas Sauerwald, Balanced allocations in batches: Simplified and generalized, 2022.
- ▶ \_\_\_\_\_, Balanced allocations with the choice of noise, 2022.
- ▶ \_\_\_\_\_, Tight bounds for repeated balls-into-bins, 2022.

# **Bibliography III**

- ▶ Dimitrios Los, Thomas Sauerwald, and John Sylvester, *The power of filling bins*, 2022.
- M. Mitzenmacher, On the analysis of randomized load balancing schemes, Theory Comput. Syst. 32 (1999), no. 3, 361–386. MR 1678304
- ▶ M. Mitzenmacher, B. Prabhakar, and D. Shah, *Load balancing with memory*, The 43rd Annual IEEE Symposium on Foundations of Computer Science, 2002. Proceedings., IEEE, 2002, pp. 799–808.
- ▶ Y. Peres, K. Talwar, and U. Wieder, Graphical balanced allocations and the  $(1 + \beta)$ -choice process, Random Structures Algorithms **47** (2015), no. 4, 760–775. MR 3418914
- ▶ M. Raab and A. Steger, "Balls into bins"—a simple and tight analysis, Proceedings of 2nd International Workshop on Randomization and Approximation Techniques in Computer Science (RANDOM'98), vol. 1518, Springer, 1998, pp. 159–170. MR 1729169